Running MLwiN from within Stata: the runmlwin command

> e-Stat meeting University of Bristol 7th April 2011

George Leckie and Chris Charlton Centre for Multilevel Modelling University of Bristol

INTRODUCTION

Existing multilevel modelling commands in Stata

- Stata provide the xtmixed, xtmelogit and xtmepoisson commands to fit multilevel models
 - Limited range of models can be specified
 - Computationally quite slow to fit models
- Sophia Rabe-Hesketh and Anders Skrondal provide the gllamm command
 - Wide range of models can be specified
 - Computationally slow to fit models
- Other user-written multilevel modelling commands include: hlm, realcomimpute, runmplus, sabre, winbugs

Multilevel modelling in MLwiN

1. Estimation of multilevel models for continuous, binary, ordered categorical, unordered categorical and count data

2. Fast estimation via classical and Bayesian methods

3. Estimation of multilevel models for cross-classified and multiple membership non-hierarchical data structures

4. Estimation of multilevel multivariate response models, multilevel spatial models, multilevel measurement error models, multilevel multiple imputation models and multilevel factor models

RAUDENBUSH (1993) CROSS-CLASSIFED MODELLING EXAMPLE

Scottish neighbourhood study on child educational attainment

- Scottish neighbourhood study on child educational attainment
- 2310 students nested within 17 schools and 524 neighbourhoods
- First analysed by Garner and Raudenbush (1991)
- Re-analysed by Rabe-Hesketh and Skrondal (2008), Raudenbush (1993), Raudenbush and Bryk (2002) and others



Journal of Educational Statistics Winter 1993, Vol. 18, No. 4, pp. 321–349

A Crossed Random Effects Model for Unbalanced Data With Applications in Cross-Sectional and Longitudinal Research

Stephen W. Raudenbush *Michigan State University*

Key words: hierarchical models, maximum likelihood, covariance components

Hierarchical linear models have found widespread application when the data have a nested structure—for example, when students are nested within classrooms (a two-level nested structure) or students are nested within classrooms and classrooms are nested within schools (a three-level nested structure). Often, however, the data will have a more complex nested structure. In Example 1, students are nested within both neighborhoods and schools; however, a school can draw students from multiple neighborhoods, and a neighborhood can send students to multiple schools. In Example 2, children are nested within classrooms during the first year of the study; however, each child finds himself or herself with a new teacher and a new set of classmates H

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TABLE 1

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Organization of data in the Scotland neighborhood study: 1	Numbers of observations in each neighborhood by school cell
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Crossed Random Effects

TABLE 3

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Descriptive statistics from the Scotland neighborhood study

Variable	п	m	SD
(a) Neighborhood level			
Social deprivation	524	0.037	0.622
(b) Student level			
Total attainment	2310	0.093	1.00
Primary 7 VRQ	2310	0.506	10.65
Primary 7 reading	2310	-0.044	13.89
Dad occupation	2310	-0.464	11.78
Dad education			
(1 = Schooling past 15; $0 = $ No $)$	2310	0.215	0.41
Mom education			
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Sex $(1 = M; 0 = F)$	2310	0.480	0.50

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TABLE 4

Modeling results for Scotland neighborhood data

(a) Fixed effects		Model 1: ariance ig	nored		Model 2: triance esti	imated	Model 3: Fitted model			
Predictor	Coefficient	SE	t ratio	Coefficient	SE	t ratio	Coefficient	SE	t ratio	
Grand mean	0.080	0.028		0.078	0.027		0.086	0.028		
Primary 7 VRQ							0.028	0.002	12.18	
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Sex							0.056	0.028	1.97	
Neighborhood										
social deprivation							-0.156	0.026	-6.03	
(b) Variance component	ts									
Parameter			Estimate			Estimate			Estimate	
Neighborhoods										
$Var(\pi_{0r}^{(1)}) = \tau_{00}$			0.200			0.141			0.006	
Schools										
$\operatorname{Var}(\pi_c^{(2)}) = \delta^2$						0.075			0.004	
Students										
$\operatorname{Var}(e_{it}) = \sigma^2$			0.804			0.800			0.454	
(c) Model fit										
Deviance						3074.26			1487.00	

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IGLS ESTIMATION

Model 1 fitted using IGLS

$$attain_{i} = \beta_{0} + u_{\text{neighid}(i)}^{(2)} + e_{i}$$
$$u_{j}^{(2)} \sim N(0, \sigma_{u(2)}^{2}), \qquad e_{i} \sim N(0, \sigma_{e}^{2})$$

. runmlwin attain cons, ///

level2(neighid: cons) ///

level1(studentid: cons)

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Model 2 fitted using IGLS

 $attain_i = \beta_0 + u_{\text{schid}(i)}^{(3)} + u_{\text{neighid}(i)}^{(2)} + e_i$

 $u_j^{(3)} \sim N(0, \sigma_{u(3)}^2), \qquad u_j^{(2)} \sim N(0, \sigma_{u(2)}^2), \qquad e_i \sim N(0, \sigma_e^2)$

- . tabulate schid, gen(s)
- . forvalues i = 1/16 {

constraint define `i' [RP3]var(s`i') = [RP3]var(s17)

• }

. runmlwin attain cons, ///

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level3(cons: s1-s17, diagonal) ///
level2(neighid: cons) ///
level1(studentid: cons) constraints(1/16)
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Model 3 fitted using IGLS

 $\begin{aligned} attain_{i} &= \beta_{0} + \beta_{1}p7vrq_{i} + \beta_{2}p7read_{i} + \beta_{3}dadocc_{i} + \beta_{4}daded_{i} \\ &+ \beta_{5}momed_{i} + \beta_{6}dadunemp_{i} + \beta_{7}male_{i} + \beta_{8}deprive_{i} \end{aligned}$

 $+u_{\text{schid}(i)}^{(3)} + u_{\text{neighid}(i)}^{(2)} + e_i$

 $u_j^{(3)} \sim N(0, \sigma_{u(3)}^2), \qquad u_j^{(2)} \sim N(0, \sigma_{u(2)}^2), \qquad e_i \sim N(0, \sigma_e^2)$

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level3(cons: s1-s17, diagonal) ///
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RP2 var(cons)	0.202	0.141	0.004	-	
RP3 var(s1) var(s2) var(s17)		0.075 0.075 0.075	0.004 0.004 0.004	_	
RP1 var(cons)	0.804	0.799	0.456	_	
Statistics deviance	6415.970	6356.711	4769.335	-	
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MCMC ESTIMATION

Model 1 fitted using MCMC

$$\begin{aligned} attain_i &= \beta_0 + u_{\text{neighid}(i)}^{(2)} + e_i \\ u_j^{(2)} &\sim \text{N}(0, \sigma_{u(2)}^2), \qquad e_i &\sim \text{N}(0, \sigma_e^2) \end{aligned}$$

```
. runmlwin attain cons, ///
```

```
level2(neighid: cons) ///
```

```
level1(studentid: cons) ///
```

```
mcmc(on) initsprevious
```

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. runmlwin attain cons, level2(neighid: cons) level1(studentid: cons) mcmc(on) initsprevious nopause >

> Number of obs 2310 =

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Normal response model Estimation algorithm: MCMC No. of Observations per Group

MLwiN 2.23 multilevel model

Level Variabl	e Groups	Minimum	Average	Maximu	m	
neighi	d 524	1	4.4	1	.6	
Burnin Chain Run time (seco Deviance (dbar Deviance (thet Effective no. Bayesian DIC) = abar) =	500 5000 5.8 6055.62 5804.62 250.99 6306.61				
attain	Mean S	Std. Dev.	z	ESS	[95% Cred.	Interval]
cons	.0831759	.0285483	2.91	1513	.0261265	.1383685
Random-effe	cts Parameters	Mean	Std. De	ev. ESS	[95% Ci	red. Int]
Level 2:	var(cons)	. 2025886	.025749	9 738	.1563706	. 2566977
Level 1:	Van(conc)	0060207	076573	5 3670	7557447	0507067

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Model 2 fitted using MCMC

$$attain_{i} = \beta_{0} + u_{\text{schid}(i)}^{(3)} + u_{\text{neighid}(i)}^{(2)} + e_{i}$$
$$u_{j}^{(3)} \sim N(0, \sigma_{u(3)}^{2}), \qquad u_{j}^{(2)} \sim N(0, \sigma_{u(2)}^{2}), \qquad e_{i} \sim N(0, \sigma_{e}^{2})$$

- matrix b = (0, .075, .15, .8)
- . runmlwin attain cons, ///

level3(schid: cons) ///

```
level2(neighid: cons) ///
```

```
level1(studentid: cons) ///
```

```
mcmc(cc) initsb(b)
```

tata/MP 11.2 - http:/	/www.stata-press.com/da	ta/mlmus2/neighbo	orhood.dta - [Res	sults]				_ 🗆 🗵
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. matrix b = (0,.075,.15,.8)							^
. runmlwin att > mc(cc) inits	ain cons, level b(b) nopause	3(schid: co	ns) level	2(neighid:	cons) le	vell(student	tid: cons)	mc
MLwiN 2.23 mul Normal respons Estimation alg	e model		N	umber of o	bs =	2310		
Level Variabl	e No. of Groups	Observa Minimum	tions per Average	Group Maximum				
schi neighi		22 1	135.9 4.4	286 16				
Burnin Chain Run time (seco Deviance (dbar Deviance (thet Effective no. Bayesian DIC) =	500 5000 6.55 6039.42 5818.77 220.65 6260.07						
attain	Mean S	td. Dev.	z	ESS [95% Cred.	Interval]		
cons	.0962458 .	0651659	1.48	228	0349854	.2170355		
Random-effe	cts Parameters	Mean	Std. De	v. ESS	[95% C	red. Int]		
Level 3:	van(conc)	0005624	040200	0 ว 175	0207661	2261022		•
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Model 3 fitted using MCMC

 $\begin{aligned} attain_{i} &= \beta_{0} + \beta_{1}p7vrq_{i} + \beta_{2}p7read_{i} + \beta_{3}dadocc_{i} + \beta_{4}daded_{i} \\ &+ \beta_{5}momed_{i} + \beta_{6}dadunemp_{i} + \beta_{7}male_{i} + \beta_{8}deprive_{i} \end{aligned}$

 $+u_{\text{schid}(i)}^{(3)}+u_{\text{neighid}(i)}^{(2)}+e_i$

 $u_j^{(3)} \sim N(0, \sigma_{u(3)}^2), \qquad u_j^{(2)} \sim N(0, \sigma_{u(2)}^2), \qquad e_i \sim N(0, \sigma_e^2)$

- . matrix b = (0, 0, 0, 0, 0, 0, 0, 0, 0, 1, 1, 1)
- . runmlwin attain cons p7vrq p7read dadocc daded ///

```
momed dadunemp male deprive, ///
level3(schid: cons) ///
level2(neighid: cons) ///
level1(studentid: cons) mcmc(cc) initsb(b)
```

Stata/MP 11.2 - http://www.stata-press.com/data/mlmus2/neighborhood.dta - [Results]	
<u>Eile Edit Data Graphics Statistics User Window H</u> elp	8
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. runmlwin attain cons p7vrq p7read dadocc daded momed dadunemp male deprive, level3(schid: > cons) level2(neighid: cons) level1(studentid: cons) mcmc(cc) initsb(b) nopause

MLwiN 2.23 multilevel model Normal response model Estimation algorithm: MCMC Number of obs = 2310

Level Variable	No. of Groups	Observa Minimum	ations per Average	Group Maximum		
schid neighid	17 524	22 1	135.9 4.4	286 16		
Burnin Chain Run time (second Deviance (dbar) Deviance (thetab Effective no. of Bayesian DIC	ar)	= 500 = 5000 = 11.4 = 4744.81 = 4704.18 = 40.63 = 4785.44				
attain	Mean	Std. Dev.	z	ESS [9	95% Cred.	Interval]
cons p7vrq p7read dadocc daded momed dadunemp male deprive	.0908623 .0275759 .0262263 .0080741 .1427545 .060508 1224706 0556571 1562529	.028849 .0022769 .0017901 .0013762 .0411568 .0379714 .0468447 .0281836 .0260999	12.11 14.65 5.87 3.47 1.59 -2.61 -1.97	4555 .0 4683 .0 4678 .0 5455 .0 47070 45032 46651)341885)232042)226469)053839)616242)130428 2136019 1103975 2078412	.1490847 .0320144 .0297636 .0107383 .2235425 .1339119 0290561 0002866 1057263

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TABLE 4

Modeling results for Scotland neighborhood data

(a) Fixed effects	Model 1: School variance ignored				Model 2: triance esti	imated	Model 3: Fitted model		
Predictor	Coefficient	SE	t ratio	Coefficient	SE	t ratio	Coefficient	SE	t ratio
Grand mean	0.080	0.028		0.078	0.027		0.086	0.028	
Primary 7 VRQ							0.028	0.002	12.18
Primary 7 reading							0.026	0.002	14.98
Dad occupation							0.008	0.001	5.94
Dad education							.0143	0.041	3.51
Mom education							0.059	0.037	1.59
Dad unemployment							-0.121	0.047	-2.56
Sex							0.056	0.028	1.97
Neighborhood									
social deprivation							-0.156	0.026	-6.03
(b) Variance component	ts								
Parameter			Estimate			Estimate			Estimate
Neighborhoods									
$Var(\pi_{0r}^{(1)}) = \tau_{00}$			0.200			0.141			0.006
Schools									
$\operatorname{Var}(\pi_c^{(2)}) = \delta^2$						0.075			0.004
Students									
$\operatorname{Var}(e_{it}) = \sigma^2$			0.804			0.800			0.454
(c) Model fit									
Deviance						3074.26			1487.00

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<pre>. estimates table model1b model2b model3b, stats(deviance) keep("FP1:" "RP2:" "RP3:" "RP1:") Variable model1b model2b model3b FP1</pre>										8
<pre>. estimates table model1b model2b model3b, stats(deviance) keep("FP1:" "RP2:" "RP3:" "RP1:") > b(%9.3f) style(oneline) Variable model1b model2b model3b FP1</pre>	📂 🖬 🖷 🗎 🖻 🕶 📶	⊾ - 🗹 • 🗹 🔂 .								_
FP1 0.083 0.096 0.035 p7vrq 0.028 0.028 p7read 0.026 0.008 dadocc 0.143 0.066 dadunemp -0.122 0.056 deprive -0.156 0.005 RP2 var(cons) 0.203 0.142 0.005 var(cons) 0.203 0.142 0.005 RP3 var(cons) 0.100 0.006 RP1 var(cons) 0.806 0.800 0.457 Statistics 0.806 0.800 0.457				el3b, stats(d	leviance)	keep("FP1:"	"RP2:"	"RP3:"	"RP1:")	
cons 0.083 0.096 0.035 p7vrq 0.028 0.028 p7read 0.026 0.028 dadocc 0.008 0.143 daded 0.143 0.061 momed -0.122 0.056 daprive -0.156 0.056 RP2 0.203 0.142 0.005 var(cons) 0.203 0.142 0.005 RP3 var(cons) 0.100 0.006 RP1 var(cons) 0.806 0.800 0.457 Statistics	Variable	model1b	model2b	model3b	- -					
var(cons) 0.203 0.142 0.005 RP3 var(cons) 0.100 0.006 RP1 var(cons) 0.806 0.800 0.457 Statistics	cons p7vrq p7read dadocc daded momed dadunemp female	0.083	0.096	0.028 0.026 0.008 0.143 0.061 -0.122 0.056						
var(cons) 0.100 0.006 RP1 var(cons) 0.806 0.800 0.457 Statistics		0.203	0.142	0.005						
var(cons) 0.806 0.800 0.457 Statistics			0.100	0.006						
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WORK EFFICIENTLY

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Bristol_e-Stat.do

```
52
53
       * Open the Raudenbush (1993) Scotish neighbourhood study data
54
       use "http://www.stata-press.com/data/mlmus2/neighborhood.dta", clear
55
56
       * Replicate Table 1 of Raudenbush (1993). The table shows that the data are
       * cross-classified. Students are said to be nested within the cross-
57
       * classification of schools by neighbourhoods. To account for both school
58
       * and neighbourhood effects in our models of student attainment, we will
59
       * need to fit cross-classified multilevel model using runmlwin.
60
61
       table neighid schid if inrange(neighid, 26, 38) | inrange(neighid, 251, 263) ///
62
           inrange(neighid,793,800)
63
64
       * Replicate Table 3 of Raudenbush (1993). The table gives summary statistics.
65
       egen pickone = tag(neighid)
66
67
       tabstat deprive if pickone==1, stat(mean sd) format(%4.3f) columns(stats)
68
69
       tabstat attain p7vrg p7read dadocc daded momed dadunemp male, ///
70
           stat(mean sd) format(%4.3f) columns(stats)
71
72
73
       * Generate a unique student identifer variable which will be the level 1
74
       * unit identifier variable in the runmlwin command
75
       gen studentid = n
76
77
       * Generate a variable cons to act as the constant or intercept variable in
78
       * the runmlwin models
79
       gen cons = 1
80
81
       * Sort the data by students within neighbourhoods (otherwise runmlwin will
82
       * complain when we try to fit the following model)
83
       sort neighid studentid
84
85
       * Fit a two-level (students within neighbourhoods) variances components
86
       * model to attain. This model is refered to as model 1 in Table 4 of
87
       * Raudenbush (1993). Note, you will need to click the "Resume Macro" button
4
```

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RESOURCES TO HELP YOU LEARN RUNMLWIN

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Advice	Contents What's New News	

help runmlwin

<u>Title</u>

runmlwin - Run the MLwiN multilevel modelling software from within Stata

<u>Syntax</u>

runmlwin responses_and_fixed_part, random_part [discrete(discrete_options)] [mcmc(mcmc_options)]
 [general_options]

where the syntax of responses_and_fixed_part is one of the following

for univariate continuous, binary, proportion and count response models

depvar indepvars [if] [in]

for univariate ordered and unordered categorical response models

depvar indepvars1 [(indepvars2, contrast(numlist)) ...] [if] [in]

where *indepvars1* are those independent variables which appear with separate coefficients in each of every log-odds contrast, while *indepvars2* are those independent variables which appear with common coefficients for those log-odds contrasts specified in **contrast(***numlist***)**. Contrasts can be thought of as the separate "subequations" or "arms" of a multinomial response model. These contrasts are indexed 1,2,... up to the total number of contrasts included in the model. The total number of contrasts will be one less than the number of response categories.

for multivariate response models

```
(depvar1 indepvars1, equation(numlist))
        (depvar2 indepvars2, equation(numlist))
        [(depvar3 indepvars3, equation(numlist))]
        [...]
        [if] [in]
```

where **equation**(*numlist*) specifies equation numbers. Equation numbers are indexed 1,2,... up to the total number of equations (i.e. response variables) included in the model.

and the syntax of *random_part* is

[...] [level2(levelvar: [varlist] [, random_part_options])] level1(levelvar: [varlist] [, random_part_options])



Examples

IMPORTANT. The following examples will only work on your computer once you have installed MLwiN and once you have told **runmlwin** what the mlwin.exe file address is. See *Remarks on installation instructions* above for more information.

(a) Continuous response models

Two-level models

Setup

. use http://www.bristol.ac.uk/cmm/media/runmlwin/tutorial, clear

Two-level random-intercept model, analogous to xtreg (fitted using IGLS) (See page 28 of the MLwiN User Manual) (You will need to click the "Resume macro" button twice in MLwiN to fit the model.) . runmlwin normexam cons standlrt, level2(school: cons) level1(student: cons)

Two-level random-intercept and random-slope (coefficient) model (fitted using IGLS) (see page 59 of the MLwin User Manual)

. runmlwin normexam cons standlrt, level2 (school: cons standlrt) level1 (student: cons)

Refit the model suppressing the two pauses in MLwiN (fitted using IGLS) (See page 59 of the MLwiN User Manual)

. runmlwin normexam cons standlrt, level2 (school: cons standlrt) level1 (student: cons) nopause

Refit the model, where this time we additionally calculate the level 2 residuals (fitted using IGLS) *(See page 59 of the MLwiN User Manual)*

. runmlwin normexam cons standlrt, level2 (school: cons standlrt, residuals(u)) level1 (student: cons)

Two-level random-intercept and random-slope (coefficient) model with a complex level 1 variance function (fitted using IGLS)

(See page 99 of the MLwin User Manual)

- . matrix A = (1,1,0,0,0,1)
- . runmlwin normexam cons standlrt girl, level2(school: cons standlrt) level1(student: cons standlrt girl, elements(A))

Two-level random-intercept and random-slope (coefficient) model using MCMC (where we first fit the model using IGLS to obtain initial values for the MCMC chains) (see page 71 of the MLwiN MCMC Manual)

- . runmlwin normexam cons standlrt, level2 (school: cons standlrt) level1 (student: cons)
- . runmlwin normexam cons standlrt, level2 (school: cons standlrt) level1 (student: cons) mcmc(on) initsprevious

Multivariate response models

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MLwiN	Kerunmlwin: Running MLwiN from within Stata	
Realcom	runmlwin is a Stata command which allows Stata users to run the powerful MLwiN multilevel modelling soft	ware from
MLPowSim	within Stata.	indie frein
runmlwin	The multilevel models fitted by runmlwin are often considerably faster than those fitted by the Stata's xtm	ixed.
Presentations	xtmelogit and xtmepoisson commands. The range of models which can be fitted by runmlwin is also mu	uch wider
Examples	than those commands. runmiwin also allows fast estimation on large data sets for many of the more comp multilevel models available through the user written gliamm command.	lex
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User Forum	MLwiN has the following features:	
CMM software support	1. Estimation of multilevel models for continuous, binary, count, ordered categorical and unordered categ	jorical data
	Fast estimation via classical and Bayesian methods	
	3. Estimation of multilevel models for cross-classified and multiple membership nonhierarchical data struct	
	 Estimation of multilevel multivariate response models, multilevel spatial models, multilevel measuremer models and multilevel multiple imputation models 	it error
	These details with a screen shot are available on our runmlwin leaflet (pdf, 0.1mb)	
	Presentations	
	We have provided a range of presentations showcasing runmlwin . These presentations provide a quick on how the command works and the range of models which can be fitted. <u>More > ></u>	overview of

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Realcom	 UK Stata Users' Group, 17th Meeting (16th September 2011) 							
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→ Examples	 University of Bristol, Mplus/MlwiN User Group (MUGS) meeting (14th June 2011) Slides (PDF, 2.3mb) 							
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CMM software support	 Modern Modeling Methods (M3) Conference, University of Connecticut (26th May 2011) 							
	 <u>Slides</u> (PDF, 3.2mb) State de file (de, 0.1mb) te replicate all applyage presented in the slides 							
	 <u>Stata do-file</u> (do, 0.1mb) to replicate all analyses presented in the slides. 							
	 2011 American Sociological Association Spring Methodology Conference, Tilburg University (20th May Slides (PDF, 2.0mb) 	/ 2011)						
	 <u>Stata do-file</u> (do, 0.1mb) to replicate all analyses presented in the slides. 							
	 University of Bristol, e-Stat meeting (7th April 2011) 							
	 <u>Slides</u> (PDF, 1.7mb) 							
	 <u>Stata do-file</u> (do, 0.1mb) to replicate all analyses presented in the slides. 							
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• 8th International Amsterdam Multilevel Conference (17th March 2011)

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CMM software support	 Rasbash, J., Charlton, C., Browne, W.J., Healy, M. and Cameron, B. 2009. MLwiN Version 2.1. Centre 1 Modelling, University of Bristol. 	for Multilevel 🔺
	For models fitted using MCMC estimation, we ask that you additionally cite:	

Browne, W.J. 2009. MCMC Estimation in MLwiN, v2.13. Centre for Multilevel Modelling, University of Bristol.

Papers using runmlwin

Please let George Leckie (<u>g.leckie@bristol.ac.uk</u>) know of any further publications using **runmlwin** including forthcoming papers, books, PhD theses, etc.

- Cheung, C., Goodman, D., Leckie, G. and Jenkins, J. (2011) <u>Understanding Contextual Effects on Externalizing</u> <u>Behaviors in Children in Out-of-home Care: Influence of Workers and Foster Families</u>. *Children and Youth Services Review*, 33, 2050-2060.
- Chung, H. and Beretvas, S.N. (2011) <u>The Impact of ignoring multiple membership data structures in multilevel</u> <u>models</u>. British Journal of Mathematical and Statistical Psychology. Forthcoming.
- Leckie, G. and Baird, J.-A. (2011) <u>Rater effects on essay scoring: A multilevel analysis of severity drift, central tendency and rater experience</u>. Journal of Educational Measurement. Forthcoming.
- Leckie, G., Pillinger, R., Jones, K. and Goldstein, H. (2011) <u>Multilevel modelling of social segregation</u>. Journal of Educational and Behavioral Statistics. Forthcoming.
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- If you use runmlwin in your work, please cite runmlwin
- Leckie, G. and Charlton, C. (2011) *runmlwin: Stata module for fitting multilevel models in the MLwiN software package*. Centre for Multilevel Modelling, University of Bristol.
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